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Convergence of speech rate in conversation predicts cooperation

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1 Abstract

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3 During conversation, interlocutors coordinate their behavior on many levels. Two distinct forms of behavioral coordination have been empirically linked with affiliation and 4 cooperation during or following face-to-face interaction: *behavior matching* and 5 interpersonal synchrony. Only the latter form constitutes behavioral entrainment 6 involving a coupling between independent oscillators. We present the first study of the 7 8 association between spontaneously occurring behavioral coordination and postinteraction economic game-play. Triads of same-sexed strangers conversed for 10 min, 9 after which each participant played an unannounced one-shot prisoner's dilemma (PD) 10 toward each co-participant. When dyads had higher language style matching scores 11 (LSM: Gonzales et al., 2010), the individuals evaluated each other more positively, but 12 they were no more likely to cooperate in the PD. However, when dyads' speech rates 13 14 (mean syllable duration) converged more strongly from the beginning to the end of the conversation, they were more likely to cooperate in the PD, despite no effect on 15 interpersonal evaluations. Speech rate convergence, a form of rhythmic entrainment, 16 could benefit interlocutors by mutually reducing cognitive processing during interaction. 17 We suggest that spontaneous, temporally-based behavioral coordination might facilitate 18 19 prosocial behavior when the joint cooperative effort is itself perceived as a form of coordination. 20 21 22 23 24 25 26 27 28 **1. Introduction** 29

1

Conversational interaction is fundamental to human communication, and involves the dynamic interplay of many complex phenomena. While engaged in conversation, interlocutors communicate with their bodies, voices, and language. Research across many disciplines has documented a variety of ways that conversationalists coordinate their actions in the service of mutually beneficial interaction. How people talk together in real time is closely tied to broader interactive goals, which themselves are products of adaptations for navigating the social world.

38 Two distinct forms of behavioral coordination have been empirically linked with 39 affiliation and cooperation during or following face-to-face interaction (Bernieri & 40 Rosenthal, 1991; Hove & Risen, 2009). The first, *behavior matching*, involves individual *B* copying a behavior of individual *A*, but with neither a particular temporal relation to 41 42 A's action, nor any implication that A responds in any specific fashion to B's copying 43 action. A substantial body of research has established that people unconsciously mimic 44 their interaction partners' postures, gestures, and mannerisms (Lakin et al., 2003), and 45 language use patterns (Niederhoffer & Pennebaker, 2002), and that such mimicry is 46 related to subsequent affiliative behavior. Among a large number of similar findings, 47 people spontaneously mimic an experimental confederate's gestures and report greater 48 liking for a confederate who mimics them (Chartrand & Bargh, 1999), and leave larger tips for a waitress who mimics them (van Baaren et al., 2003). Researchers using the 49 50 automated Linguistic Inquiry and Word Count algorithm (Pennebaker et al., 2001, 2007) 51 have found that similarity in relative usage frequency of common function word 52 categories (e.g. prepositions, conjunctions) predicts successful hostage negotiations

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53 (Taylor & Thomas, 2008), task group cohesiveness (Gonzales et al., 2010), and the formation and persistence of romantic relationships (Ireland et al., 2011). Coordinated 54 55 language use and behavior may facilitate mutual understanding (Pickering & Garrod, 2004). Ireland and Pennebaker (2010; see also Meyer & Bock, 1999) argued that function 56 57 words such as pronouns and articles (unlike content words) are "inherently social," 58 because their comprehension typically depends, not just on the conventions of a speech 59 community, but on shared frames of reference actively established among interlocutors. 60 For example, every English-speaker knows the meaning of *garden*, but the particular 61 garden referred to by *the garden* will be apparent to a listener only when she shares the same immediate frame of reference as the speaker. For this reason, according to Ireland 62 and Pennebaker (2010), pairwise similarity in frequency of function word use is 63 64 associated with greater affiliation or cooperation.

65 A second form of behavioral coordination is *interpersonal synchrony*, which typically involves entrainment—a temporal coupling between independent oscillators that 66 67 enter into some type of phase relationship. Prime examples of this are turn-taking in conversation (Wilson & Wilson, 2005) and playing music with an isochronous beat 68 69 (Bispham, 2006). In Wilson and Wilson's (2005) model of conversational turn-taking, 70 speech rate entrainment occurs via speakers' syllabic production, which operates 71 interpersonally as a medium for entraining neural oscillators among interlocutors. This facilitates conversational coordination and allows for inter-turn transitions marked by 72 73 minimal gap and minimal overlap (Stivers et al., 2009). Perceptions of timing in music and speech can affect subsequent productions in these respective domains (Jungers et al., 74

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2002), and speech rate convergence has been linked to interpersonal judgments (e.g.
ratings of competence: Street, 1984).

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78 Talk is just one form of social interaction in which people are sensitive to entrainment. Studies have shown that singing together can increase cooperation in a 79 80 prisoner's dilemma game (Anshel & Kipper, 1988) and a public goods game (Wiltermuth 81 & Heath, 2009), though the effect can be sensitive to experimental conditions (e.g. 82 Kurzban, 2001). Children who sang and danced together were more likely to assist one 83 another in a later playground incident (Kirschner & Tomasello, 2010). Synchronous tapping, but not asynchronous tapping, generated higher affiliation ratings, but only when 84 85 the synchrony was with another person, and not just experienced (i.e., tapping to a 86 metronome) (Hove & Risen, 2009). Synchronized training in competitive rowers resulted 87 in increased endorphin release (Cohen et al., 2010), suggesting a proximate mechanism 88 motivating this kind of behavioral coordination. Behavioral entrainment is highly detectable, and can impact people's perceptions of the affiliation between the 89 synchronizers. Hagen and Bryant (2003) showed that better temporal coordination in a 90 91 music performance positively affected third party judgments of coalition quality between 92 the musicians. While social entrainment may have evolved in many species from the 93 simpler adaptive ability to entrain one's behavior to rhythmic information in the physical 94 environment (Phillips-Silver et al., 2010), human interpersonal synchrony is moderated by many social factors and interacts in complex ways with group membership and the 95 96 dynamics of alliance formation (Miles, Griffiths et al., 2009; Miles, Lumsden, et al., 97 2011).

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99	Laughter is another interactive phenomenon that can involve behavioral
100	coordination and may be associated with cooperative behavior. Research has shown that
101	people who have known each other longer tend to laugh together more (Bryant, 2012;
102	Smoski & Bachorowski, 2003a) and familiarity between conversationalists is perceptible
103	in the co-laughter itself (Bryant, 2012). Lynch (2011) found that people with greater
104	similarity in implicit preferences laugh together more, suggesting an association with
105	social cohesion. Gervais and Wilson (2005) argued that laughter functions as a medium
106	for mirthful emotional contagion that recruits partners into resource-building social play.
107	Accordingly, comparative work has demonstrated that chimpanzees use laugh-like
108	vocalizations to manage playful social interactions, and that antiphonal laugh sequences
109	lengthen play time (Davila-Ross et al., 2011). Other scholars have suggested a variety of
110	communicative functions for coordinated laughter that relate to cooperation (Owren &
111	Bachorowski, 2003; Mehu & Dunbar, 2008), bonding (Dezecache & Dunbar, 2012;
112	Platow et al., 2005) and social assortment (Flamson, et al., 2011).
113	
114	The adaptive significance of these various phenomena remains a matter of debate.
115	Simple mimicry in nonhuman social animals has obvious adaptive advantages (e.g.
116	treating conspecifics' fear responses as reliable cues of imminent danger), and is
117	presumably the phylogenetic source of more elaborate forms of behavioral coordination

118 (Lakin et al., 2003). However, why these should serve as "social glue" is unclear. A

119 number of non-human animal species exhibit inter-individual temporal coordination (Hall

120 & Magrath, 2007), but the functions of these displays often remain unknown. Phillips-

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Silver et al. (2010) argue that even in cognitively simple species, collective social 121 122 entrainment can amplify social signals in adaptive ways (e.g. courtship choruses; 123 Greenfield, 1994). In human collective action, social entrainment may be necessary to accomplish work activities that require behavioral coordination. Recent work has shown 124 125 that engaging in synchronized action facilitates success in later joint activity. For 126 example, people who rocked synchronously in chairs, compared to controls that rocked 127 asynchronously, were better able to subsequently coordinate their action on a 128 collaborative task (Valdesolo et al., 2010). This suggests that synchronizing action may 129 calibrate expectations about others' behavior, and help coordinate action in other domains. 130

131

132 In this study, we examined whether distinct kinds of vocal and verbal convergence 133 in naturalistic social interactions predicted cooperation in a one-shot prisoner's dilemma 134 (PD). In a PD, an actor chooses whether to cooperate or defect toward a recipient. The actor gains the largest payoff when he defects while the recipient cooperates; the second 135 136 largest when both cooperate; the third largest when both defect; and the lowest when the 137 actor cooperates while the recipient defects. From a strictly monetary perspective, 138 defection is always the best decision in a one-shot PD. However, a sizeable proportion of 139 educated American, European, and Japanese participants treat one-shot PDs as assurance 140 games, gaining the most psychological utility from mutual cooperation (Fehr & Camerer, 141 2007; Hayashi et al., 1999; Kiyonari et al., 2000), and therefore cooperating if, and only 142 if, they expect their partner to cooperate. This suggests that social preferences transform 143 the PD into a coordination game (specifically, a Stag Hunt – Van Huyck et al., 1990) in

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which one coordinated outcome (mutual cooperation) yields higher payoffs to bothplayers than the other coordinated outcome (mutual defection).

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147 To assess whether different types of naturally occurring behavioral coordination 148 facilitate cooperation-as-coordination, we measured behavior among strangers in open-149 ended conversation prior to their playing an unannounced one-shot simultaneous PD. We examined dyadic convergence in three vocal characteristics: (1) fundamental frequency 150 151 (F_0) ; (2) variation in F_0 , and (3) speech rate (mean syllable duration). We also calculated 152 several measures of coordinated laughter and laughter/speech coordination. Finally, we 153 calculated each dyad's language style matching score (LSM: Gonzales et al., 2010). We 154 also examined the relationships between convergence and coordination in these diverse channels. Based on the empirical literature reviewed above, we expected that greater 155 behavioral convergence would raise expectations of cooperative coordination, and that 156 157 therefore dyads showing greater (1) vocal convergence, (2) coordinated laughter and (3) verbal convergence (higher LSM score) would be more likely to cooperate in the PD. We 158 also elicited ratings of co-participants' warmth and competence, and predicted that these 159 160 person perception variables would mediate the relationship between the 161 convergence/coordination variables and PD decisions. This is the first study to examine 162 whether spontaneous (as distinct from experimentally induced) behavioral coordination is 163 associated with post-interaction behavior in an incentivized social dilemma.

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165	The analyses presented here build on our previous report of findings regarding the
166	determinants of our conversation participants' PD decisions (Gervais et al., 2013). In a
167	multivariate model, we found two main effects: people were more likely to cooperate (1)
168	if they grew up in a wealthier zip code and (2) towards more facially attractive co-
169	participants. We also found two interaction effects with subclinical primary psychopathy
170	(callous affect, interpersonal manipulation) as measured by a self-report instrument:
171	people who scored higher on primary psychopathy were less likely to cooperate toward
172	co-participants (1) who interrupted them more frequently during the pregame
173	conversation, and (2) with whom they discovered no "common ground" (i.e. reliable cues
174	to future interaction). This model explained 15.6% of the variance in probability of
175	cooperating. Our goal in the present research is to assess which, if any, measures of
176	verbal or vocal convergence improve the predictive power of this model.

178 2. Methods

More detailed descriptions of (1) the participant pools, (2) the conversation and post-conversation game-play and questionnaire procedures, (3) the attractiveness rating procedure, and (4) the conversation transcription procedures can be found in Gervais et al. (2013).

183 2.1 Participants

184 Conversation participants (*n* = 105) were undergraduates at UCLA. All
185 participants were given a \$10 show-up fee; 90% of them were also fulfilling a course
186 requirement. Participants were all native speakers of English, their median age was 19

187 years, and their ethnic composition corresponded closely to that of the multi-ethnic

188 campus population. Conversation groups were same-sex triads (20 female, 15 male).

189 2.2. Procedures

Conversation participants were grouped equidistantly around a small circular table. After determining that the conversation participants were strangers to each other, an experimenter recited a prepared script asking the participants to converse for 10 minutes on any topic(s) they wished. Participants were informed that their conversation would be videotaped, but were given no details about the post-conversation procedure. Conversations were recorded using a Canon Vixia HV30 camcorder (Audio: MP2 compression, 384 kbps) connected to an Audio-Technica U841a omnidirectional

197 condenser boundary microphone (30 Hz – 20 kHz frequency response).

198 Following the conversation, participants sat at visually isolated laptops running z-Tree version 2.1 (Fischbacher, 2007). First, participants played a one-shot PD toward 199 200 each of their two co-participants. Choices were labeled "Keep" \$3 provided by the 201 experimenter (=defect) or "Transfer" the \$3 to the co-participant, whereupon it would be 202 doubled to \$6 (=cooperate). To ensure the confidentiality of participants' PD choices, 203 they were instructed, truthfully, that one of the three of them would not receive their 204 earned payoff, but instead a randomly generated but realistic set payoff. Participants then 205 rated each of their co-participants on "warmth" and "competence" using separate sliders, 206 completed a well-validated self-report psychopathy instrument (the LSRP: Levenson et 207 al., 1995), and answered a set of basic demographic questions (age, ethnicity, childhood 208 zip code).

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210 2.3 Data analysis

211	Perceived Warmth and Competence. Participants' ratings of their co-participants'
212	<i>warmth</i> and <i>competence</i> were moderately to strongly congruent (Cronbach's $\alpha = 0.69$).
213	Therefore, we averaged the standardized <i>warmth</i> and <i>competence</i> ratings of each
214	participant towards each co-participant. We refer to this measure as positive person
215	perception (PPP). For all significant results incorporating PPP, separate analyses using
216	warmth and competence produced qualitatively very similar results.
217	Language Style Matching (LSM). Because of the large time and training
218	investment required for transcriptions and data analyses reported elsewhere (Gervais et
219	al., 2013), a portion of each 10-minute conversation was pre-selected for transcription
220	and further analysis. This portion included the first 60 seconds of the conversation and
221	two other sections of \geq 60 seconds duration from minutes 2-5 and 6-10 of the
222	conversation. Start times of the second and third transcription periods were moved back
223	toward the beginning of the video, if necessary, so that all transcription periods began
224	with the start of a new conversational topic. Total time transcribed per group ranged from
225	3.02-5.57 minutes ($M \pm SD = 4.08 \pm 0.68$ min). For the LSM analyses, we constrained the
226	transcriptions to yield only English words spelled as indicated in the Linguistic Inquiry
227	and Word Count (LIWC) 2007 program dictionary (Pennebaker et al., 2007). The LIWC
228	algorithm calculates, for a sample of speech or text, the proportion of words in a text that
229	fall into each of 67 categories, not all of which are mutually exclusive. For theoretical
230	reasons reviewed above, Pennebaker and colleagues (Gonzales et al., 2010; Ireland &

231 Pennebaker, 2010; Ireland et al., 2011; Pennebaker et al., 2003) have placed particular emphasis on interpersonal similarity in the usage frequency of nine types of function 232 233 words: auxiliary verbs (e.g. *am*, *will*, *have*), articles, common adverbs (e.g. *hardly*, *often*), personal pronouns, indefinite pronouns, prepositions, negations, conjunctions and 234 235 quantifiers.

236 We ran the LIWC algorithm on each participant's speech output during the 237 transcribed portions of the conversation. To determine overall language style 238 convergence within dyads, we first calculated the correlations between co-participants' usage (i.e. the percentage of each individual's total words uttered) for each of the nine 239 240 function word categories.

241 Following Gonzales et al. (2010), we next calculated each co-participant dyad's LSM score based on inter-individual similarities in the proportions of the nine types of 242 243 function words. The LSM score of a dyad, Person 1 and Person 2, with regard to a 244 particular function word type, e.g. quantifiers, is calculated as:

245 quanLSM = 1 - (|quan1 - quan2|/(quan1 + quan2))

where quan1 is the percentage of Person 1's words that are quantifiers, and quan2 is the 246 247 percentage of Person 2's words that are quantifiers. An LSM score can range from 0.00-248 1.00. Each dyad's total LSM is calculated as the mean of its LSM scores across the nine 249 categories of function words.

250 To determine whether dyads with higher LSM scores were more likely to 251 cooperate in the PD, we ran a logistic regression with PD decision (cooperate or defect) 252 as the dependent variable. To account for the non-independence of each individual's two 11

253 PD decisions, we calculated robust standard errors of the regression coefficients,

clustering by individual, before calculating confidence intervals and *p*-values.

255 Vocal Characteristics. For the vocal analyses, we also divided the 10-minute 256 conversations into three sections (not the same sections used for LSM analyses). Section 1 was 0:00-3:20, Section 2 was 3:21-6:40, and Section 3 was 6:41-10:00. Audio files 257 were exported from the video recordings using Apple iMovie software and saved as 44.1 258 259 kHz, 16-bit way files. Using Adobe Audition 3, we then extracted the first continuous 260 five-second portion of continuous speech for each speaker and for each section (i.e., three clips per participant) that did not contain overlapping speech, other than cases of one-261 262 word backchannels (e.g., uhhuh) or other vocal noises. Of the 315 clips (105 participants \times 3 time periods), 10 contained no speech excerpts that met these criteria; these were 263 264 deleted case-wise in subsequent statistical analyses.

265 The extracted clips were analyzed acoustically using Praat, version 5.3.01 266 (Boersma & Weenink, 2011). We measured mean fundamental frequency (F_0) (the 267 acoustic correlate of perceived pitch), fundamental frequency standard deviation (F_0 SD) (acoustic correlate of perceived pitch variability) and mean syllable duration (MSD) 268 269 (speech rate) for each clip. F_0 was measured using the autocorrelation method in Praat 270 with default pitch settings suggested by Praat for men (100-500 Hz) and women (120-600 271 Hz). Octave jump errors and other analytical errors, such as F_0 estimates during voiceless 272 segments, were fixed through pitch setting adjustment (never exceeding +/- 20 Hz adjustment in the lower limit, and +/- 60 Hz in the higher limit), or removed manually. In 273 274 cases where small overlapping vocalizations occurred in the extracted clips, the 275 overlapped portions were removed prior to analysis. On average, >90% of the original

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clips were analyzed, with most requiring no editing. MSD was calculated by dividing the
total time of speech energy determined through visual analysis of the spectrogram in the
clip by the number of spoken syllables (i.e., not written word syllables). (Spectrogram:
FFT method, Gaussian window shape, dynamic range - 50 dB).

To determine whether co-participants generally converged with respect to F_0 , F_0 SD, and MSD, we treated each dyad as a data point. For each of these variables in each conversation section, we regressed the value of one member of the dyad on the value of the other member of the dyad. Significantly positive slopes indicate greater than chance similarity within dyads. To determine whether co-participants became more similar in these variables over time, we used Wald tests to compare slopes across conversation sections.

287 To test whether convergence in vocal characteristics affected PD play, we first calculated, for every co-participant dyad in each of the three conversation sections, the 288 289 absolute value of the difference between their values for each variable. For each dyad, we 290 then estimated the slope (β) of the linear regression line formed by the three points (section 1, section 2 and section 3). Negative slopes indicate decreasing differences (i.e. 291 292 increasing similarity) over time between the two co-participants with regard to that 293 variable. We then ran logistic regressions in which each dyad was a data point, PD 294 decision (cooperate or defect) was the dependent variable and the relevant β value 295 (standardized) was the independent variable. Significantly negative relationships indicate 296 that vocal convergence increases the probability of cooperation.

13

297 *Laughter Analyses.* We coded laughter throughout the 10-min duration of every conversation. Two coders counted laugh instances in all conversations using video 298 playback. A bout of laughter was defined as a series of nonverbal, vocalized calls often 299 with successive expiratory elements, though sometimes containing only a single call. We 300 included voiced (i.e., tonal) and unvoiced bouts. Laughter is typically produced with an 301 302 initial burst amplitude and frequency that decays over time (Titze et al., 2008). Laugh 303 calls judged as a single bout had to originate from the same initial burst. Bouts that were 304 back-to-back without a noticeable pause, as evidenced by perceptible re-initialized energy, were counted as one laugh. Laugh counts across speaker conditions were highly 305 correlated between the two coders (Cronbach's alpha = 0.96) so data from one coder were 306 307 used in our analyses. We calculated the values of an individual-level variable (raw 308 number of laughs produced) and four dyad-level variables: (1) raw number of simultaneous laughs (co-laughter count), (2) simultaneous laughs divided by the sum of 309 310 the two dyad members' total laughs (co-laughter proportion); (3) laughs by the first dyad member while the second dyad member was speaking (laughs during other's speech) and 311 312 (4) summed laughs while the other dyad member was speaking, divided by the sum of the 313 two dyad members' total laughs (bi-directional laughter during other's speech).

We used Akaike's Information Criterion (AIC: Akaike, 1974) to assess the effects of adding independent variables to models that successfully predicted our dependent variables. AIC takes into account the tradeoff between a predictive model's accuracy (which should be maximized) and its complexity, or number of independent variables (which should be minimized). In a comparison between two models, the one with the

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lower AIC value is better, as it more closely approximates the causal processes thatgenerated the data.

321

322 **3. Results**

323

324 3.1 Language style matching

325

326 Co-participants generally matched their language styles with respect to function words. Table 1 shows correlation coefficients (Pearson *r*) of co-participants' usage 327 328 percentages of the nine function word categories. For only two of these categories (conjunctions and quantifiers), co-participant pairs failed to attain highly significant 329 330 similarity. Because we found, consistent with other research (e.g. Newman et al., 2008), 331 some sex differences in function word use (e.g. compared to men, women used more 332 auxiliary verbs [13.6% vs. 11.9%, Cohen's d = 0.58, p = 0.004]), we also ran the 333 correlation analyses separately for the two sexes. Among women, co-participant pairs failed to attain significant similarity only for prepositions, conjunctions and quantifiers; 334 among men, co-participant pairs failed to attain significant similarity only for articles, 335 336 conjunctions and quantifiers. All other within-sex co-participant correlations were 337 significant at p < 0.01. Among the 105 dyads, the mean LSM score was 0.82 (SD = 0.08). 338 339

339 Co-participant dyads that were sampled for longer periods of time had higher 340 LSM scores ($\beta = 0.025 \pm 0.011$, n = 105, p = 0.032), as would be expected if longer 341 sampling periods reduce error variance, i.e. the impact on LSM of random intra-

individual variation in function word use. We therefore calculated the residuals of this
regression (i.e. LSM score relative to amount of time sampled) and used these values as
an independent variable to confirm results obtained using raw LSM scores as the
independent variable.

346

347	LSM scores were not significantly associated with Prisoner's Dilemma decisions.
348	Bivariate logistic regressions revealed non-significant trends towards more likely
349	defection given higher LSM scores, which is opposite to that predicted (using raw LSM
350	scores: logistic regression with standard errors based on clustering by decision-maker's
351	identity, odds ratio \pm SE = 0.062 \pm 0.132, <i>n</i> = 206, <i>p</i> > 0.1; using residuals on time
352	sampled: o.r. \pm SE = 0.159 \pm 0.333, <i>n</i> = 206, <i>p</i> > 0.3). When we added either raw LSM
353	scores or residual scores to the multivariate predictive model described in Gervais et al.
354	(2013), (1) neither variable had an independent significant relationship with PD
355	decisions, (2) the resulting models did not increase the proportion of variance explained,
356	and (3) they increased the AIC (Akaike, 1974). Thus, even after taking into account all
357	known effects of independent variables on PD decisions by our participants, LSM scores
358	had no explanatory value with respect to predicting PD decisions.
250	

359

360

361 3.2. Vocal characteristics

362

363 Table 2 shows the results (β coefficients with standard errors) of regressing, for 364 each of the three acoustic variables (F_0 , F_0 SD, and mean syllable duration) in each

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365 conversation section, each participant's value on one of his or her co-participants' value. 366 That is, each data point is a dyad of co-participants. For male F_0 , these coefficients were 367 significantly negative, i.e. if one male of a dyad had a high F_0 , his co-participants tended 368 to have low F_0 values, at both the beginnings and the ends of conversations. For female 369 F_0 , and for F_0 SD in both sexes, there was no relationship between co-participants' values. 370 There were no significant changes between β values from section 1 to section 3.

371

However, mean syllable duration did show inter-individual convergence over the course of the conversations. In section 1, the regression coefficient was non-significantly negative, whereas in section 3, it was significantly positive and, therefore, significantly different from the β of section 1. Furthermore, during section 2, the β value was intermediate between sections 1 and 3 (0.043 ± 0.092) and not significantly different from either.

378

379 For F_0 and F_0 SD, we found no relationship between inter-individual convergence 380 (i.e. the slope over time of the absolute value of the difference between co-participants' 381 values) and probability of cooperating in the Prisoner's Dilemma. Indeed, for F_0 , there 382 was a marginally significant trend for higher slopes (i.e. greater inter-individual 383 differentiation over time) to be associated with cooperation (logistic regression with 384 standard errors based on clustering by decision-maker's identity, o.r. \pm SE = 1.46 \pm 0.30, 385 n = 178, p = 0.061). For F_0 SD, there was no relationship between convergence and PD 386 decision (o.r. \pm SE = 1.33 \pm 0.30, *n* = 172, *p* > 0.20). However, in dyads that converged 387 more strongly in mean syllable duration, participants were more likely to cooperate in the

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PD (o.r. \pm SE = 0.57 \pm 1.40, *n* = 176, *p* = 0.02). Of the three vocal variables, only MSD convergence improved the predictive power of the multivariate model described in Gervais et al. (2013). When added to this model, MSD convergence had a significant (*p* = 0.04) independent effect on the probability of cooperating, and adding MSD convergence to the model increased r² from 0.156 to 0.196, and decreased AIC from 228.7 to 190.6, indicating a closer approximation of the causal processes that generated the data.

395 3.3. Laughter

396 Across all 35 conversations, over 1000 laughs were counted in total ($M \pm SD =$ 397 29.1 ± 13.0), and a substantial percentage of these were produced in coordination (i.e., co-laughter) with at least one other speaker (M \pm SD = 41.4% \pm 16.8%). Women 398 produced more laughs than men ($n_{\text{female}} = 60$, $n_{\text{male}} = 45$, $M_{\text{female}} \pm \text{SD} = 17.2 \pm 7.7$, $M_{\text{male}} =$ 399 400 12.3 ± 8.8 , d = 0.59, p = 0.003). Among dyads (n = 206 PD decisions in all analyses), 401 neither co-laughter count (o.r. \pm SE = 1.03 \pm 0.05, p > 0.50), co-laughter proportion (o.r. \pm SE = 1.41 \pm 1.71, *p* > 0.50), laughs during other's speech (o.r. \pm SE = 1.00 \pm 0.06, *p* > 402 0.50), nor bi-directional laughter during other's speech (o.r. \pm SE = 1.56 \pm 1.86, p > 403 404 0.50) was associated with the probability of cooperating in the prisoner's dilemma. When added to the multivariate model predicting PD play described by Gervais et al. (2013), 405 406 none of these independent variables had a significant independent effect on PD play, and 407 all of them increased AIC. Although we made no predictions about sex differences in the 408 relationships between behavioral convergence and PD play, a post-hoc analysis showed 409 that only among men (n = 90 PD decisions), dyads with higher co-laughter counts (o.r. \pm 410 SE = 1.14 ± 0.07 , p = 0.03) and co-laughter proportions (o.r. \pm SE = 26.77 ± 42.53 , p =

411	0.04) were significantly more likely to cooperate. Running the Gervais et al. (2013)
412	multivariate model separately for men and women revealed that in men only, AIC was
413	reduced by adding co-laugher count or co-laughter proportion as an independent variable.
414	Neither variable had a significant independent effect on PD play in men, but the effect of
415	co-laughter proportion approached significance at $p = 0.08$.
416	
417	3.4. Associations among independent variables
418	
419	LSM score was not associated with any of the vocal or laughter variables, nor
420	were F_0 or F_0 SD associated with any of the laughter variables. However, dyads that
421	converged more with respect to MSD (i.e. had more strongly negative slopes) had higher
422	co-laughter counts ($r = -0.22$, $p = 0.04$).
423	
424	3.5 Positive Person Perception
425	
426	Higher PPP ratings were marginally associated with an increased probability of
427	cooperating in the PD (odds ratio \pm SE = 1.43 \pm 0.29, <i>n</i> = 204, <i>p</i> = 0.07).
428	LSM scores were significantly associated in the predicted (positive) direction with
429	participants' positive person perception (PPP) ratings. This held whether the independent
430	variable was raw LSM score (linear regression with robust standard errors clustered by
431	participant: β = 1.93 ± 0.76, <i>N</i> = 208, <i>p</i> = 0.013) or residual of LSM on time sampled (β =
432	2.33 \pm 0.84, <i>N</i> = 208, <i>p</i> = 0.007). However, LSM scores did not predict behavior in the
433	PD (see above).

434 None of the vocal characteristics, including MSD, significantly predicted positive 435 person perception scores ($F_0: \beta = -0.55 \pm 0.60$, N = 180, p > 0.30; F_0 SD: $\beta = 0.22 \pm 0.53$, 436 N = 174, p > 0.50; MSD: $\beta = 0.44 \pm 4.53$, N = 178, p > 0.50).

Only one laughter variable, laughs during other's speech, was associated with positive person perception. PPP ratings were higher in dyads in which the two participants laughed more while the other person was speaking ($\beta = 0.97 \pm 0.43$, N = 208, p = 0.027). However, laughs during other's speech did not predict behavior in the PDG (see above). Neither co-laughter count ($\beta = 0.019 \pm 0.014$, N = 208, p = 0.20) nor colaughter proportion ($\beta = 0.44 \pm 0.42$, N = 208, p > 0.20) was associated with PPP ratings.

444 4. Discussion

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446 We examined the relationships between vocal and verbal convergence in a 447 spontaneous conversation and the participants' subsequent decisions in a Prisoner's Dilemma game. Existing empirical work in communication led us to predict that behavior 448 449 matching in language use and vocal convergence in prosodic features of speech would be 450 associated with cooperation in a PD game. One form of vocal convergence (speech rate) 451 was positively related to the probability of PD cooperation. However, although we found 452 strong evidence for language style matching in zero-acquaintance small groups 453 (consistent with others' findings; Niederhoffer & Pennebaker, 2002; Newman et al., 454 2008), LSM was unrelated to post-conversation PD decisions. The LSM results are 455 surprising in view of other work (Gonzales et al., 2010; Ireland & Pennebaker, 2010; 456 Ireland et al., 2011) suggesting that cooperation in several contexts (e.g. group task

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performance, romantic relationships, even long-term scholarly collaborations) can be
predicted using the same language style matching (LSM) metric (Gonzales et al., 2010)
that we applied to our data.

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One important difference between earlier LSM research and the current study 461 462 was that we did not cue the importance of substantive cooperation among our participants 463 until after the conversation. We told them only that we were studying "small talk among 464 strangers," and that they would be answering some questions at the end of the 465 conversation. In contrast, the experimental situation of a task group (Gonzales et al., 466 2010), and the real-life situations of a speed-date (Ireland et al., 2011) or a hostage negotiation (Taylor & Thomas, 2008), presumably foreground the detection of 467 cooperative potential in one's interlocutor(s) before the face-to-face interaction. One 468 469 possibility is that, with respect specifically to function word use, close style matching 470 (i.e. LSM scores >0.75) is the typical outcome of collaborative conversation, and pairwise style *divergence* follows from declines in affiliation or trust in potentially 471 472 agonistic situations. In contrast, our study's experimental conversation context was 473 friendly and collaborative, with very little (apparently) at stake. Therefore, style matching occurred (Table 1), but LSM was unrelated to subsequent, and unanticipated, PD 474 475 decisions. This is consistent with the relationship of LSM to PPP even in the absence of 476 an effect of LSM on cooperation.

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478 Co-participants did tend to cooperate more as a function of how much their479 speech rates converged. Behavioral convergence that involves entrainment (i.e.,

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temporally based) might be distinct from other forms of convergence (e.g., behavior 480 matching) because it provides immediate mutual benefits. For example, becoming 481 482 entrained can introduce synergy that could potentially reduce mutual metabolic costs of interacting (Marsh et al., 2006). This would make entrainment a form of coordination, in 483 484 which profitable cheating is impossible, but players' uncertainty about each other's 485 choices may still yield suboptimal outcomes (Van Huyck et al., 1990). Successful 486 behavioral entrainment may reduce uncertainty in future coordination by indexing how 487 well co-participants can coordinate their action. For example, speech rate entrainment 488 may be a reliable indicator that an interlocutor can coordinate his or her actions with 489 one's own actions in a rapid and fine-tuned manner, mutually reducing the cognitive processing costs of interaction. Given that our participants may typically view a one-shot 490 491 PD as a coordination game (Fehr & Camerer, 2007; Hayashi et al., 1999; Kiyonari et al., 492 2000), coordination in speech rate may increase perceived ability to coordinate on 493 cooperation in a PD, raising rates of cooperation. This is consistent with the effect of speech rate convergence on cooperation even in the absence of an effect on positive 494 495 person perception – the perceptions of coordination that lead to cooperation do not 496 necessarily require positive interpersonal evaluations. DeSteno et al. (2012) likewise 497 found that disengagement gestures performed by a humanoid robot reduced participants' 498 donations and expectations of donations in a social dilemma, yet did not affect 499 participants' liking of the robot. Future research should use multi-dimensional person 500 and relationship perception measures to tap the relevant attributions and evaluations that 501 underlay perceived coordination capacity.

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503 We found that language matching had no impact on cooperative decisions in the 504 PD game, whereas speech rate entrainment did increase the probability of cooperation. It 505 may be that behavior matching is more subject to vigilance against cheating than is synchrony because matching is more used in manipulation (Dawkins & Krebs, 1981). 506 507 Coordination is mutually beneficial and offers no incentive for defection, while mimicry 508 and other unilateral forms of behavior matching are potentially intentional and 509 manipulative (Bailenson et al., 2008; Bourhis et al., 1975). Pardo et al. (2010) found that 510 when conversationalists were instructed to imitate one another covertly, they often 511 converged phonetically (a form of behavior matching), but simultaneously diverged in 512 articulation rates (a form of entrainment). This suggests greater success at manipulative matching than entrainment. Bailenson et al. (2008) found that mimicry had negative 513 impacts on trustworthiness and warmth judgments when it was explicitly noticed— 514 515 suggesting a sensitivity to manipulation—whereas even instructed, consciously mediated 516 synchrony (e.g., intentionally walking in time, clapping together, or swinging a cup while singing) can enhance cooperation despite explicit awareness of the behavioral 517 convergence (Valdesolo & DeSteno, 2011; Valdesolo et al., 2010; Wiltermuth & Heath, 518 519 2009). Our results fit this pattern, even though language style matching may be less likely 520 than gestural or postural mimicry to be consciously detected, and conscious attempts to 521 match others' language styles are generally unsuccessful (Ireland & Pennebaker, 2010). 522 523

We did not find a relationship between coordinated dyadic laughter and game play across all participants, either in the absolute amount of laughing in response to another person, or in the proportion of all laughter in a triad shared by a dyad within it. We did,

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however, discover an unexpected sex difference. The more a male dyad laughed together, 526 527 the more likely they were to cooperate in the PD game. Women laughed significantly 528 more than men, a finding consistent with other studies of laughter in small groups of strangers (Bryant 2012; Smoski & Bachorowski, 2003a), but women's laughter was not 529 related to game play. This suggests the intriguing possibility that male co-laughter in 530 531 zero-acquaintance contexts has relatively higher cue validity for cooperative intentions 532 and/or the ability to coordinate in the future. Kurzban (2001) found that low level social 533 signals such as mutual eye gaze, gentle touching, as well as instant virtual messages, 534 increased cooperation relative to a control condition between men but not between women in a public goods game. The tendency of men, but not women, to cooperate more 535 in response to simple social cues might reflect a difference in the forms and functions of 536 537 intra-sexual coalitions (e.g., Hess & Hagen, 2006; Rucas et al., 2010; Vigil 2007).

539 Laughter between established friends, however, does not quite follow the apparent pattern for strangers. Research on laughter in developing friendships revealed that 540 541 antiphonal laughter (i.e., sequential call and response laughter) occurred earlier in 542 women's friendships than in men's, and was established at least three weeks into the relationship, as opposed to males who took up to six weeks to increase antiphonal laugh 543 544 frequency (Smoski & Bachorowski, 2003b). Laughter between conversationalists not 545 only increases in frequency as people become friends, but also in form. Bryant (2012) 546 found several acoustic differences in laughter between friends and strangers, and that 547 third parties could detect friendship from very brief (< 2 s) instances of co-laughter. 548 Laughter signals clearly play an important role in social interaction, and the functions of

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interlocutors laughing together vary depending on relationship context, social strategies,and group composition (Bryant & Aktipis, in review).

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The current research illustrates how studies of conversational behavior can inform 552 553 work on the evolution of cooperation. A limitation of our study is that we traded off 554 experimental control for ecological validity—we therefore cannot document a causal 555 relationship, but we found that some forms of conversational coordination were 556 associated with cooperative behavior in a naturalistic interaction. Future research should 557 vary the protocol described here by cueing the importance of both cooperation and 558 competition before the conversation, without revealing the post-conversation social dilemma. In addition, researchers should explore the perception of affiliation between 559 560 those engaged in conversation and investigate the possibility that some of these 561 coordinated behaviors are designed to transmit coalition information. Finally, these 562 results are based on the behavior and social interactions of American undergraduates, a subpopulation where many are living away from family and established social networks, 563 and therefore possibly more interested in establishing new friendships with strangers. 564 565 Further research should explore the cross-cultural validity of these findings, especially in 566 relatively closed societies where social ties are longer in duration, and social mobility is 567 lower. The dynamics of conversation can reveal a great deal about how people interact on 568 many levels, and much work remains.

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